

A Brief Guide to Causal Inference in Social Science Research

Frost Center Colloquium

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March 18, 2026

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been...”*

The essential question of causality



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2003 NLCS, Game 6. Cubs lead 3–0 in the 8th
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The Cubs surrender 8 runs and lose the series.



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The Cubs surrender 8 runs and lose the series.

*What would have happened if he hadn't
touched it?*



Today's Agenda

Talk ~30 min

- 1 Motivation: causation vs. correlation
- 2 The fundamental problem of causal inference
- 3 When RCTs are not feasible
- 4 Three quasi-experimental methods:
 - ▶ Difference-in-Differences (DiD)
 - ▶ Regression Discontinuity (RD)
 - ▶ Synthetic Controls
- 5 Further reading

Workshop ~20 min

- Share your research questions
- Identify which method(s) might apply
- Open Q&A

Come ready to talk about your work!

A Thought Experiment

Imagine waking up tomorrow as a member of the **U.S. Food and Drug Administration**.

Your first task: assess the effectiveness of a new blood pressure medication, **Potentia™**, in lowering patient blood pressure.

Your approach: release the drug publicly and compare blood pressure levels of those who *take it* with those who *do not*.

Is this a good approach?

Why or why not? What could go wrong with this comparison?



“What might have been?”

- Did the policy *cause* the outcome, or would it have happened anyway?
- Does attending college *cause* higher earnings, or do higher-ability individuals both attend college *and* earn more?
- Does moving to a better neighborhood *cause* better child outcomes, or do motivated families both move *and* have children who are observably better off?

The core problem

We observe the world only once. We cannot observe what would have happened without the policy, treatment, or event — the **counterfactual**.

Potential Outcomes and Selection Bias

Notation (Rubin, 1974):

- $Y_i(1)$: outcome for unit i if treated
- $Y_i(0)$: outcome for unit i if not treated
- We only ever observe **one** of these — the *fundamental problem of causal inference*

Average Treatment Effect on the Treated (ATT):

$$\text{ATT} = \mathbb{E}[Y_i(1) - Y_i(0) \mid D_i = 1]$$

Selection bias

A naive comparison $\mathbb{E}[Y \mid D=1] - \mathbb{E}[Y \mid D=0]$ *conflates* the treatment effect with **selection bias** — treated and untreated units are systematically different.

The Gold Standard: Randomized Controlled Trials

Random assignment eliminates selection bias:

$$\{Y_i(0), Y_i(1)\} \perp\!\!\!\perp D_i$$

Treatment is independent of potential outcomes, so a simple difference in means recovers the ATE — no confounding.

Basic requirements for a good RCT:

- Clear, well-defined treatment
- Verified balance across groups
- Pre-registered analysis plan
- IRB approval and informed consent

Moving to Opportunity (HUD, 1994)

Randomly assigned housing vouchers to low-income families.

Chetty, Hendren & Katz (2016) linked MTO data to IRS tax records and found large positive effects on children's long-run earnings — *especially* for those who moved young.

When RCTs Are Not Feasible

Random experiments are the gold standard... but often infeasible:

Cost

Large-scale randomization is expensive and logistically demanding

Ethics

We cannot randomly assign poverty, incarceration, or environmental exposure

Real-world policy

Policies are already enacted; we want to evaluate what actually happened

Solution: Find settings where treatment is *as good as random*
— **quasi-experimental methods**

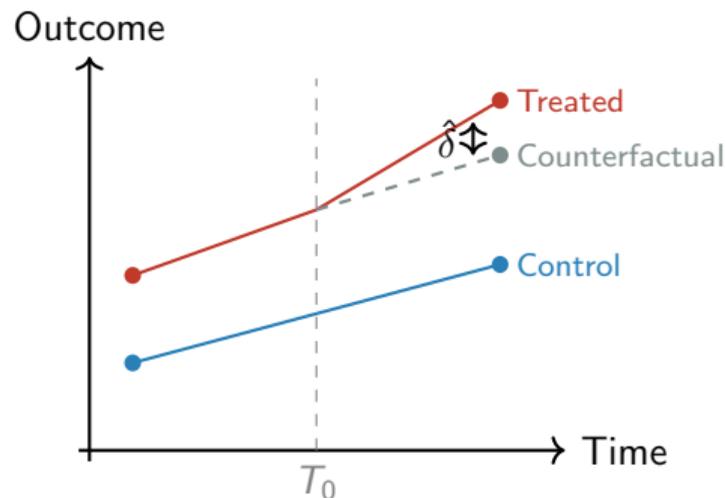
Difference-in-Differences: The Idea

Setting:

- Panel data: units observed before and after an event
- Some units receive treatment at time T_0 ; others do not
- Compare the *change* in outcomes for treated units to the *change* for untreated units

Classic example: Card & Krueger (1994) — New Jersey raised its minimum wage; neighboring Pennsylvania did not. Did NJ restaurant employment fall?

The control group provides the **counterfactual trend**.



Difference-in-Differences: The Math

DiD estimator (2×2 case):

$$\hat{\delta}_{\text{DiD}} = \underbrace{(\bar{Y}_{T,\text{post}} - \bar{Y}_{T,\text{pre}})}_{\text{treated group change}} - \underbrace{(\bar{Y}_{C,\text{post}} - \bar{Y}_{C,\text{pre}})}_{\text{control group change}}$$

Equivalent regression:

$$Y_{it} = \alpha + \beta \text{Treated}_i + \gamma \text{Post}_t + \delta (\text{Treated}_i \times \text{Post}_t) + \varepsilon_{it}$$

$\hat{\delta}$ is the ATT estimate — the coefficient on the interaction term.

Two-way fixed effects (TWFE) — multiple periods and units:

$$Y_{it} = \alpha_i + \lambda_t + \delta D_{it} + \varepsilon_{it}$$

α_i absorbs unit-level differences; λ_t absorbs common features across units within time unit.

Parallel Trends Assumption

In the *absence of treatment*, the treated and control groups would have followed the **same trend** over time:

$$\mathbb{E}[Y_{it}(0) - Y_{i,t-1}(0) \mid \text{Treated}_i = 1] = \mathbb{E}[Y_{it}(0) - Y_{i,t-1}(0) \mid \text{Treated}_i = 0]$$

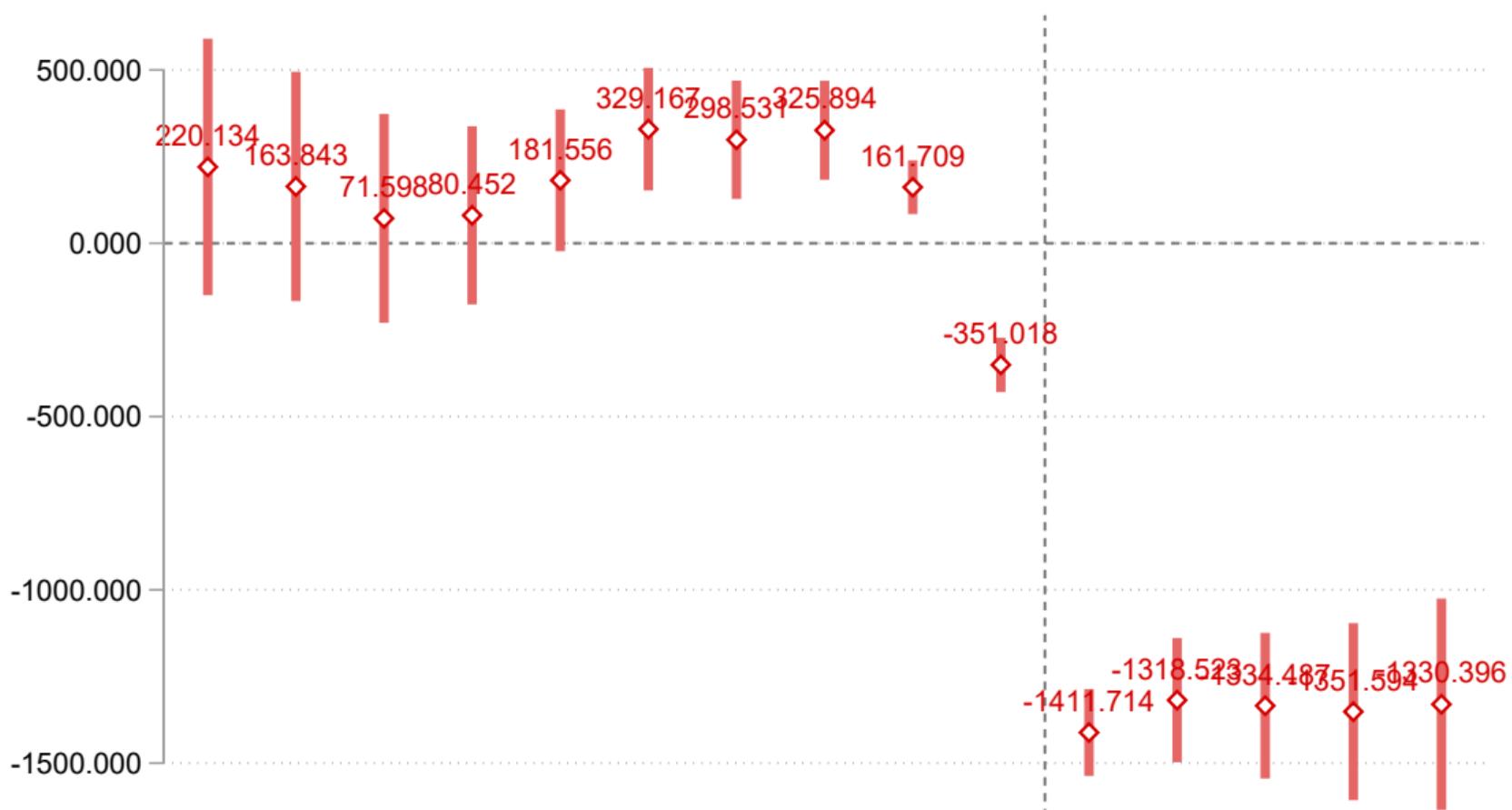
How to support the assumption:

- Plot pre-treatment outcome trends
- Event study: test whether pre-period “effects” are near zero
- Choose a plausible, similar comparison group

Common threats:

- Treated units are fundamentally different
- Simultaneous policy changes in treated units
- Anticipation effects before T_0
- Differential time trends

Admissions to State Prison



Simple 2×2 OLS:

```
# treated:post interaction = delta_DiD
model <- lm(
  outcome ~ treated * post,
  data = panel_data
)
summary(model)
```

Two-way fixed effects (TWFE):

```
library(fixest)

# Unit + time fixed effects
twfe <- feols(
  outcome ~ i(year, treated, ref = 0) |
    unit_id + year,
  data = panel_data,
  cluster = ~unit_id
)

# Event study plot
iplot(twfe, main = "Event Study")
```

fixest (Bergé, 2018) is fast and handles clustering, multiple fixed effects, and event-study plots in one workflow.

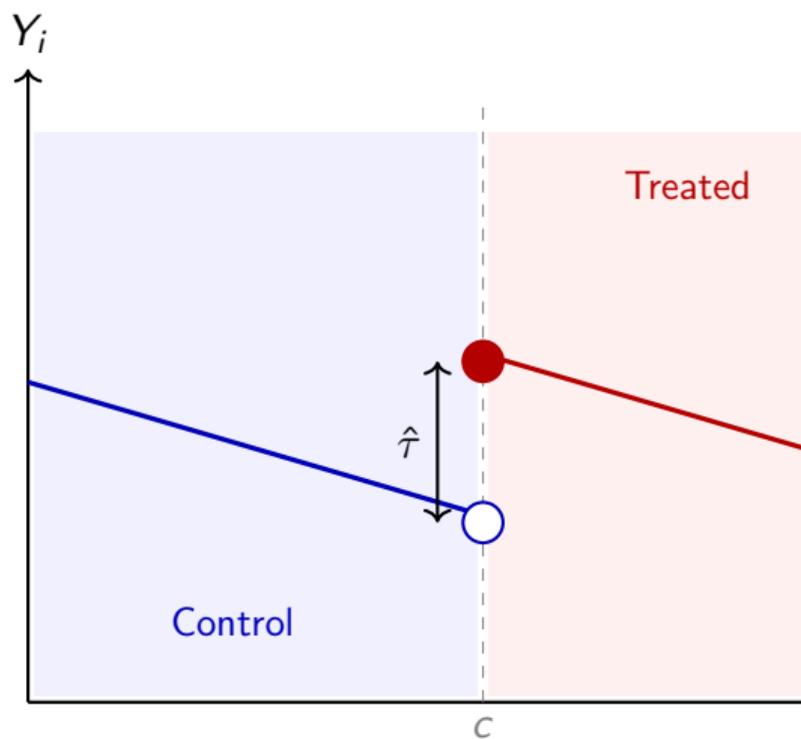
Regression Discontinuity: The Idea

Setting: Treatment is determined by whether a **running variable** X_i crosses a known **cutoff** c .

- Units just above and just below c are *very similar*
- Any discontinuous jump in the outcome at c is the causal effect of treatment

Classic examples:

- Test score \geq cutoff \Rightarrow scholarship eligibility
- Vote share $\geq 50\%$ \Rightarrow election winner
- Age $\geq 65 \Rightarrow$ Medicare eligibility



Regression Discontinuity: The Math

Sharp RD: Treatment is a deterministic function of X_i :

$$D_i = \mathbf{1}[X_i \geq c]$$

RD estimand — Local Average Treatment Effect (LATE) at the cutoff:

$$\hat{\tau}_{\text{RD}} = \lim_{x \downarrow c} \mathbb{E}[Y_i | X_i = x] - \lim_{x \uparrow c} \mathbb{E}[Y_i | X_i = x]$$

Estimated with local polynomial regression on either side of c .

Key assumption: Continuity

Potential outcomes are continuous at the cutoff.

Units cannot *precisely* manipulate X_i to sort just above or just below c .

Fuzzy RD

When crossing c raises the *probability* of treatment (not a certainty), use 2SLS with $\mathbf{1}[X_i \geq c]$ as an instrument for D_i .

Regression Discontinuity: R Code

```
library(rdrobust)

# Sharp RD estimate at cutoff c
rd <- rdrobust(
  y = data$outcome,
  x = data$running_var,
  c = cutoff_value      # e.g., c = 0 or c = 50
)

summary(rd)
# Reports: local polynomial estimate, MSE-optimal bandwidth,
#          bias-corrected confidence interval

# Bin-scatter plot around the cutoff
rdplot(
  y      = data$outcome,
  x      = data$running_var,
  c      = cutoff_value,
  title  = "Regression Discontinuity"
)
```

rdrobust (Calonico, Cattaneo & Titiunik, 2014) selects the MSE-optimal bandwidth and applies bias correction automatically.

Synthetic Controls: The Quantitative Case Study

Setting:

- *One* (or very few) treated unit(s) — a state, country, or city
- Treatment is a major policy or event
- DiD is hard: no obvious comparison group; one treated region

Key idea: Construct a **synthetic** version of the treated unit from a weighted combination of *donor pool* units that best matches its pre-treatment trajectory.

Classic example: Abadie, Diamond & Hainmueller (2010) — Effect of California's 1988 tobacco tax (Prop 99) on cigarette consumption.

Why this is powerful

- Transparent: weights are explicit
- No global parallel-trends assumption needed
- Rich pre-period fit *validates* the counterfactual
- Designed for the “ $n = 1$ treated unit” problem

Synthetic Controls: Setup

Notation:

- $J+1$ units, T time periods; unit 1 is treated starting at T_0
- Units $2, \dots, J+1$ form the **donor pool** (never treated)
- \mathbf{X}_1 : vector of pre-treatment predictors for the treated unit
- \mathbf{X}_0 : matrix of pre-treatment predictors for donor units

The **synthetic control** is a convex combination of donor units:

$$\hat{Y}_{1t}(0) = \sum_{j=2}^{J+1} w_j^* Y_{jt}, \quad t > T_0$$

with weights $W^* = (w_2^*, \dots, w_{J+1}^*)$ satisfying $w_j \geq 0$ and $\sum_j w_j = 1$.

Estimated treatment effect:

$$\hat{\alpha}_{1t} = Y_{1t} - \hat{Y}_{1t}(0), \quad t > T_0$$

Synthetic Controls: Finding the Weights

Choose W^* to minimize pre-treatment discrepancy:

$$W^* = \arg \min_W \|\mathbf{X}_1 - \mathbf{X}_0 W\|_{\mathbf{V}}^2 = \arg \min_W \sum_{k=1}^K v_k \left(X_{1k} - \sum_{j=2}^{J+1} w_j X_{jk} \right)^2$$

subject to: $w_j \geq 0 \quad \forall j$ and $\sum_j w_j = 1$

What goes in \mathbf{X} ?

- Pre-treatment outcome lags (crucial)
- Economic / demographic predictors
- More pre-periods \Rightarrow more credibility

Diagnosing fit

Compare Y_{1t} to $\hat{Y}_{1t}(0)$ in the pre-period.
Close tracking before T_0 validates the post-period counterfactual.

Synthetic Controls: Inference via Placebo Tests

Traditional standard errors do not apply when $n = 1$. Instead: **permutation / placebo inference**.

- 1 Apply the synthetic control method to *each donor unit* as if it were the treated unit
- 2 Compute a placebo treatment effect for each
- 3 Compare the treated unit's estimated effect to the distribution of placebo effects

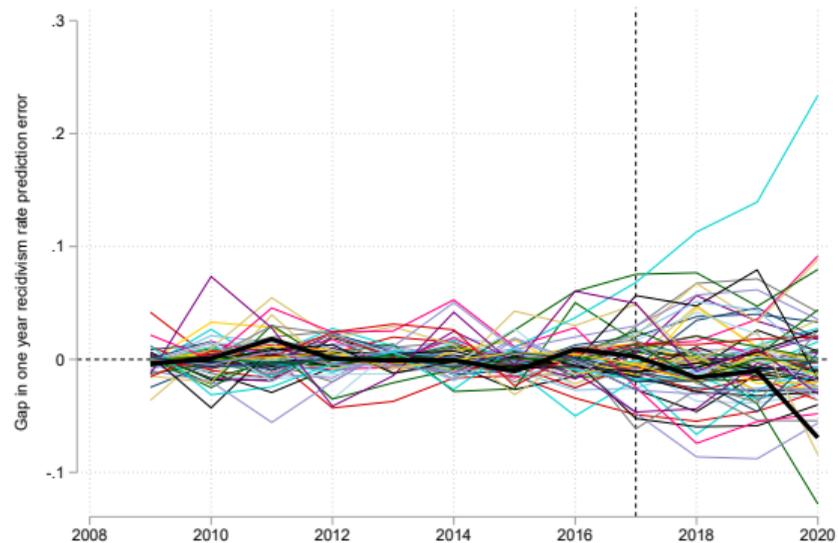
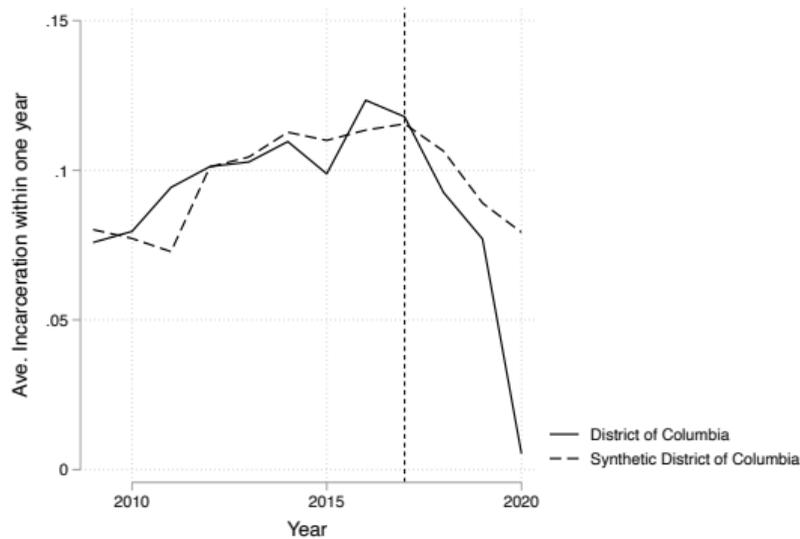
Inference rule

If the treated unit's gap is *extreme* relative to the placebo distribution, we have evidence of a real effect.

$$\text{Pseudo } p\text{-value} = \frac{\text{rank of treated unit's effect}}{\text{total number of units}}$$

Exposition follows Cunningham (2021), *Causal Inference: The Mixtape* — mixtape.scunning.com

Example: Fair Chance Housing in Washington DC



Suggestive, but not significant evidence that FCH in DC led to reduction in recidivism probabilities.

Synthetic Controls: California Prop 99

Abadie, Diamond & Hainmueller (2010): Did California's 1988 tobacco tax reduce cigarette consumption?

- **Treated unit:** California
- **Donor pool:** 38 other U.S. states with no major tobacco legislation
- **Outcome:** Per-capita cigarette sales (packs/year)
- **Synthetic California:** weighted combination of Colorado, Connecticut, Montana, Nevada, Utah, and others
- **Finding:** By 1997, California consumed ≈ 26 fewer packs/person/year than its synthetic counterfactual

Pre-treatment fit

Synthetic California closely tracks actual California from 1970–1988, lending credibility to the post-1988 counterfactual.

Placebo test

California's post-1988 gap is among the largest of any state in the donor pool — a compelling permutation test result.

Synthetic Controls: R Code

```
library(tidysynth) # modern tidy interface; install.packages("tidysynth")

synth_out <- smoking_data |>
  synthetic_control(
    outcome = cigsale, # per-capita cigarette sales
    unit = state, # unit identifier
    time = year, # time identifier
    i_unit = "California", # treated unit
    i_time = 1988, # treatment year
    generate_placebos = TRUE # enable permutation inference
  ) |>
  generate_predictor(time_window = 1980:1988,
    avg_sales = mean(cigsale),
    avg_beer = mean(beer)) |>
  generate_weights(optimization_window = 1970:1988) |>
  generate_control()

synth_out |> plot_trends() # treated vs. synthetic control
synth_out |> plot_placebos() # permutation / placebo inference
synth_out |> grab_unit_weights() # inspect donor pool weights
```

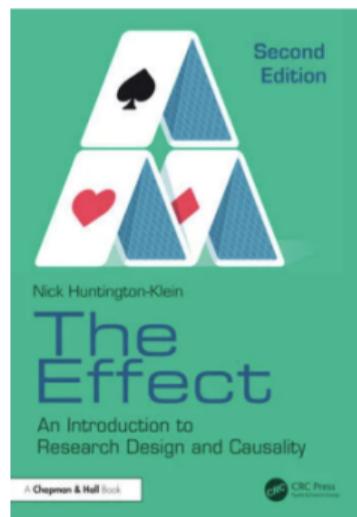
Package: *tidysynth* (Dunford, 2021). Data: *smoking* dataset shipped with the *Synth* package.

Which Method Should You Use?

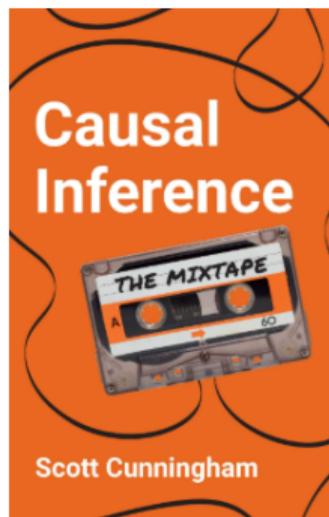
Method	What you need	Key assumption
Diff-in-Diff (DiD)	Panel data; pre/post periods; treated and control group	Parallel trends: groups would have trended together absent treatment
Reg. Discontinuity (RD)	Running variable with a known cutoff that determines treatment	Continuity at cutoff; units cannot precisely manipulate the running variable
Synthetic Controls	Aggregate units; few treated; long pre-treatment history	Pre-treatment fit is good; no interference across units

The best method is the one that most credibly mimics random assignment for your specific research setting.

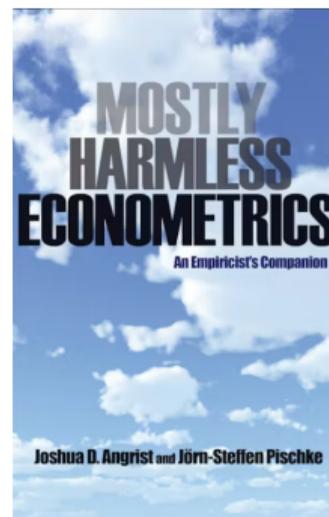
Further Reading



The Effect
Nick Huntington-Klein
Introductory; undergrad
Free online:
theeffectbook.net



**Causal Inference:
The Mixtape**
Scott Cunningham
Adv. UG / Master's
Free online:
mixtape.scunning.com



**Mostly Harmless
Econometrics**
Angrist & Pischke
Graduate / practitioner
Classic reference text

Workshop: Let's Talk About Your Research

~20 minutes — bring your research questions!

Questions to think about:

- What is your core research question?
- What is the *treatment* or *event* you want to study?
- What is your outcome variable?
- Do you have data on multiple units over time?
- Is there a natural threshold or cutoff?
- Did a policy change affect only some units or regions?

Quick method-matching guide

DiD — before/after period + treated/control group

RD — a score or threshold determines who gets treatment

Synthetic Controls — one region / country, long history, no clean comparison group

No question too basic — ask anything about methods, data, or interpretation!